# Bayesian AR model with Laplace innovations. Results and further proposals.



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Modelling time series

$$Y = (Y_1, Y_2, \dots, Y_T) \tag{1}$$

Assumptions:

- All quantities have density.
- Relation between data is described by model f(Y<sub>t</sub>|Θ, F<sub>t-1</sub>) and time-independent parameters Θ

$$f(Y_t|\Theta, \mathcal{F}_{t-1}) \Leftarrow Y_{t+1} = \Theta \Phi_t + \Sigma W_{t+1}$$
(2)

• Prior distribution  $f(\Theta|\mathcal{F}_0)$  describes beliefs - conjugate form.

In a model with constant parameters, we use Bayesian data update

$$f(\Theta|Y_t, \mathcal{F}_{t-1}) = \frac{f(Y_t|\Theta, \mathcal{F}_{t-1})f(\Theta|\mathcal{F}_{t-1})}{\int_{\Omega} f(Y_t|\Theta, \mathcal{F}_{t-1})f(\Theta|\mathcal{F}_{t-1})d\Theta}$$
(3)

In a model with Gaussian innovations and GiW (NiG) prior the estimation has two important properties

 Exponential form of density – transforms multiplication into summation of exponents

• Quadratic (polynomial) form in the exponent conserves form when summed with other quadratic form (polynomial of same order)

in a model with Laplace innovations and conjugate prior the first property holds, while the second one fails

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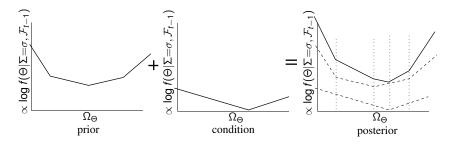


Figure: The principal of update of logarithmized conditional probability density of parameters  $f(\Theta|\Sigma, \mathcal{F}_{t-1})$  in a model with Laplace noise via Bayes rule in one dimension. The logarithm transforms multiplication into addition.

and Metropolis-Hastings or

$$J_{t} = \sum_{P_{j} \in \mathcal{C}} \frac{\operatorname{vol}(P_{j}) \Gamma(\nu_{t} - n - 1)}{2^{\nu_{t}} n!} \sum_{k=1}^{n+1} \frac{1}{H_{k}^{\nu_{t} - n - 1}} \prod_{l=1, k \neq l}^{n+1} \frac{1}{H_{l} - H_{k}}$$
(4)

where  $H_k$  is value of exponent at point  $P_k^0 \in C$ .

#### TABLE I

#### EMPIRICAL VARIANCES OF SOME ALTERNATIVE LOCATION ESTIMATORS<sup>a</sup>

(Sample Size 20)

		Di				
Estimators	Normal	10% 3 <i>a</i> <sup>b</sup>	10% 10σ <sup>c</sup>	Laplace	Cauchy	
Mean	1.00	1.88	11.54	2.10	12,548.0	Gauss
10% trimmed mean	1.06	1.31	1.46	1.60	7.3	
25% trimmed mean	1.20	1.41	1.47	1.33	3.1	
Median	1.50	1.70	1.80	1.37	2.9	Laplace
Gastwirth <sup>a</sup>	1.23	1.45	1.51	1.35	3.1	
Trimean <sup>e</sup>	1.15	1.37	1.48	1.43	3.9	

<sup>a</sup> Abstracted from Exhibit 5 in Andrews, et al. [3]. <sup>b</sup> Gaussian Mixture:  $.9\Phi(1)+.1\Phi(3)$ . <sup>c</sup> Gaussian Mixture:  $.9\Phi(1)+.1\Phi(10)$ . <sup>d</sup>  $\beta = .3\beta^{er}(1/3)+.4\beta^{e}(1/2)+.3\beta^{e}(2/3)$ , where  $\beta^{*}(\theta)$  is the  $\theta$ th sample quantile. <sup>e</sup>  $\beta = 1/4\beta^{e}(1/4)+1/2\beta^{e}(1/2)+1/4\beta^{e}(3/4)$ .

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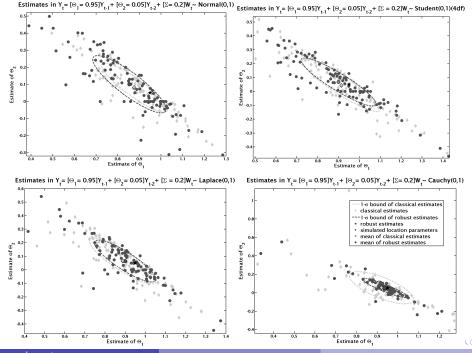


Table: Descriptive statistics of maximum likelihood estimates of location parameters in model  $y_t = (\alpha_1 = 0.95)y_{t-1} + (\alpha_2 = 0.05)y_{t-2} + 0.2e_t$  for different noises and two expected models.

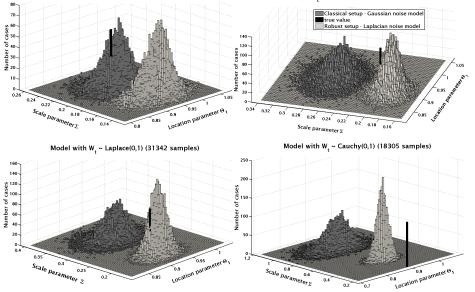
model\real		<i>N</i> (0, 1)	<i>t</i> <sub>4</sub> (0, 1)			
Inodel (real	mean $(\hat{\alpha_1}, \hat{\alpha_2})$	$cov(\hat{\alpha}_1, \hat{\alpha}_2)$	mean $(\hat{\alpha}_1, \hat{\alpha}_2)$	$\operatorname{cov}(\hat{\alpha}_1,\hat{\alpha}_2)$		
Gaussian	( .8771 )	( .0229 –.0196 )	(.9136)	( .03500322 )		
	( .0820 /	(−.0196 .0208 )	.0448	│		
Laplacian	( .8535 )	( .02850247 )	( .8988 )	( .03210281 )		
	(.1011)	(−.0247 .0281 )	( .0682 /	<u> </u>		

model\real		<i>L</i> (0, 1)	$\mathcal{C}(0,1)$			
model\rear	mean $(\hat{\alpha_1}, \hat{\alpha_2})$	$cov(\hat{\alpha}_1, \hat{\alpha}_2)$	$mean(\hat{\alpha}_1, \hat{\alpha}_2)$	$\operatorname{cov}(\hat{\alpha}_1, \hat{\alpha}_2)$		
Gaussian	( .8876 )	( .0275 –.0247 )	( .9207 )	( .02360193 )		
	.0661 /	(−.0247 .0290 )	( .0550 <i>)</i>	│		
Laplacian	( .8998 )	( .02320205 )	( .9354 )	( .01210106 )		
	( .0652 /	(−.0205 .0239 )	( .0532 )	(−.0106 .0110 )		

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Model with W, ~ Student(4df)(0,1) (33617 samples)



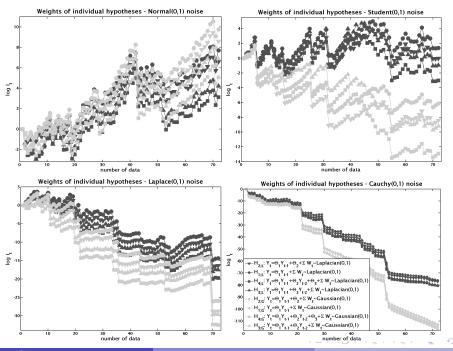
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$$\begin{array}{rcl} H_{1}: Y_{t} &=& \Theta_{1} Y_{t-1} + \sigma W_{t} \\ H_{2}: Y_{t} &=& \Theta_{1} Y_{t-1} + \Theta_{2} + \sigma W_{t} \\ H_{3}: Y_{t} &=& \Theta_{1} Y_{t-1} + \Theta_{2} Y_{t-2} + \sigma W_{t} \\ H_{4}: Y_{t} &=& \Theta_{1} Y_{t-1} + \Theta_{2} Y_{t-2} + \Theta_{3} + \sigma W_{t} \end{array}$$

$$R: Y_t = [\Theta_1 = 0.9] Y_{t-1} + [\Sigma = 0.2] W_t$$

real\hypo	$H_1+W_t\sim\mathcal{L}$	$H_2+W_t\sim\mathcal{L}$	$H_3+W_t\sim\mathcal{L}$	$H_4+W_t\sim\mathcal{L}$	$H_1 + W_t \sim \mathcal{N}$	$H_2 + W_t \sim \mathcal{N}$	$H_3 + W_t \sim \mathcal{N}$	$H_4+W_t\sim\mathcal{N}$
$\mathcal{N}^{a}(0,1)$	0.486	0.013	0.164	0.004	0.29	0.001	0.043	5.3·10 <sup>-7</sup>
$\mathcal{N}^{b}(0,1)$	0.191	0.007	0.006	0.002	0.63	0.013	0.093	0.002
$t_4(0,1)$	0.88	0.026	0.092	0.003	5.4·10 <sup>-6</sup>	3.4·10 <sup>-9</sup>	2.8·10 <sup>-8</sup>	1.5.10 <sup>-11</sup>
$\mathcal{L}(0,1)$	0.876	0.048	0.073	0.004	2.6·10 <sup>-7</sup>	3·10 <sup>-10</sup>	9.1.10 <sup>-11</sup>	7.9.10 <sup>-13</sup>
C(0,1)	0.87	0.096	0.031	0.004	7.0.10 <sup>-18</sup>	8.1.10 <sup>-21</sup>	4.5·10 <sup>-20</sup>	5.8·10 <sup>-23</sup>

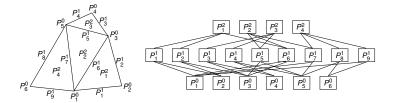
<sup>a,b</sup> The choice of hypotheses for data from model with Normally distributed noise depends on the chosen prior distribution. We therefore give results for a prior specified by density  $f(\Lambda|\mathcal{F}_0)$  the result is given in line  $\mathcal{N}^a(0, 1)$  and also a result with prior specified by  $f(\Lambda|\mathcal{F}_1)$  (posterior after one update) the result given in line  $\mathcal{N}^b(0, 1)$ .



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### Complexity:



$$N_n(r,t) = \sum_{p=n-r}^n \binom{p}{n-r} \binom{t}{p}$$

 $\lim_{t\to\infty} N_n(r,t) < t^n \quad \forall r \in \{0,1,\ldots,n\}$ 

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## Thank you for your attention.



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